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# Measuring the violence prevention climate:

## Development and evaluation of the VPC-14

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# Measuring the violence prevention climate: Development and evaluation of the VPC-14

## 1. BACKGROUND

### 1.1. Introduction

Violence and aggression are common in inpatient mental health settings (Iozzino et al., 2015) and cause problems for staff, patients and organisations (Bowers et al., 2011). Surveys suggest that 30% to 54% of health care staff have experienced violence in the past year (Campbell et al., 2011, HatchMaillette et al., 2007), and most staff working in mental health services expect to be assaulted at some time in their career (Bilgin and Buzlu, 2006). The practices that staff sometimes use to prevent violence by restricting patient autonomy (e.g. restraint, seclusion and forced medication) cause physical and psychological harm to staff and patients (Bonner et al., 2002, Fish and Hatton, 2017, Renwick et al., 2016) and can even result in patient death (Duxbury et al., 2011). In the UK, violence towards NHS staff is estimated to cost at least £69 million a year (National Institute for Health and Care Excellence, 2015a), and in mental health settings the estimated annual cost of verbal abuse and physical assaults is £20.5 million per year (Flood et al., 2008, National Institute for Health and Care Excellence, 2015b). The prevention of violence and aggression is therefore a vital part of the role of healthcare staff in these settings.

From a public health perspective, primary violence prevention refers to actions taken before violence has occurred, aimed at stopping its occurrence; secondary prevention comprises the actions taken to prevent imminent violence; while tertiary prevention encompasses the interventions that take place whilst violence is occurring and in its aftermath to minimise harm (Paterson et al., 2004). As such, primary prevention is encapsulated in activities and behaviours that are somewhat distal antecedents. There is a clear parallel, therefore, between primary violence prevention and the notion of the 'ward climate' that has been proposed as an important contextual determinant of outcomes.

The World Health Organisation (1953 p.17) stated that 'the single most important factor in the efficacy of treatment given in a mental hospital appears ... to be an intangible element which can only be described as its atmosphere'. Subsequently, the notion of social climate, the shared perceptions that people have about a particular environment (Bennett, 2010), to describe similar phenomena across social contexts including healthcare, education, and the wider workplace has become commonplace. In healthcare, the Ward Atmosphere Scale (WAS; Moos, 1974) has been the most widely used instrument but it lacks robust psychometric properties (Røssberg and Friis, 2003). As a result of the perceived shortcomings of the WAS, Schalast et al. (2008) developed the Essen Climate Evaluation Scale (EssenCES) to measure the ward atmosphere specifically in a forensic mental health setting.

## 1.2. Study rationale

The concept of a distinct 'violence prevention climate' is not entirely novel; Spector et al. (2007), who examined safety in a range of workplace settings including hospitals, describe it from an organisational perspective as employees' perceptions of the policies, procedures and training related to violence prevention rather than as an aspect of the social environment. Our review (Hallett et al., 2014) identified only one scale that measured perceptions of the violence prevention climate within mental health inpatient settings, the E13 (Björkdahl et al., 2013). However, the E13 has various shortcomings in the development and testing of the scale, specifically that there was no expert review of the items nor pilot testing prior to use, and no validity or reliability testing is described. The E13 was developed to evaluate a specific training model within one setting, and therefore may have limited generalizability away from that setting.

## 1.3. Study aims

The aim was to develop a reliable and valid scale to measure perceptions of the violence prevention climate among staff and patients in mental health inpatient settings, based on four key principles (Morello et al., 2013, Schalast et al., 2008, Streiner and Norman, 2008): i) it should be based

on available evidence on the topic e.g., violence prevention in mental health settings; ii) scale items should refer to observable and relevant phenomena, for this scale that constitute primary and/or secondary violence prevention behaviours; iii) it should provide robust and valid quantitative data that can be used to make comparisons across wards and over time; and iv) it should be quick and easy to complete by both staff and patients.

## 2. METHODS

### 2.1. Design

The study was conducted in three phases: i) scale development, ii) pilot testing, and iii) psychometric evaluation. The first phase included qualitative semi-structured interviews and focus groups, while subsequent phases utilised a cross-sectional survey design. The current study adhered to the methods of assessing reliability and validity described in the COSMIN checklist (Mokkink et al., 2010) which was designed for the evaluation of health-related patient-reported outcomes measures but are also relevant to outcomes measures more generally.

### 2.2. Participants and setting

Each phase of the study was conducted at one or more of three hospitals in England run by an independent sector not-for-profit provider of secure mental health care. Eligible participants were patients and staff residing in/working on wards in the adult male and female mental health care pathways (phases 1-3). The inclusion criteria for patient participants were age ( $\geq 18$  years), willing and able to give informed consent to participate in the study, currently admitted to inpatient mental health services, and English language speakers. Patients' clinical teams advised on whether each individual had capacity to consent to the study, and this was monitored by the researcher during the consenting process and data collection. For the focus groups in phase 1, patients were recruited via the patient experience team. Staff members of the organisation's prevention and management of violence and

aggression (PMVA) team, as well as ward-based staff, were eligible for the interviews in phase 1; they were purposively sampled to capture the view of staff who were experienced in violence prevention. In phases 2 and 3 eligible staff were those permanently employed in the clinical setting, having worked on the ward for a minimum two week period, or those employed to work in the clinical setting on a non-regular basis and having a self-expressed knowledge of the ward setting.

## 2.3. Ethical considerations

Ethical approval for all aspects of the study including instrument development and piloting was obtained from Nottinghamshire NHS NRES Research Ethics Committee, (REC reference 13/EM/0221), and the University of Northampton REC. All potential participants were provided with full information about the study. At all stages, patient participants were required to provide written informed consent. For study elements requiring only staff questionnaire completion, consent was assumed from completion and return of measures.

## 2.4. Validating instruments

Two instruments previously used in research in mental health inpatient settings were used as part of the scale validation process.

### 2.4.1.1. *EssenCES (Schalast et al., 2008)*

The EssenCES is a 15-item scale comprising three factors: therapeutic hold (perceptions of how supportive the environment is to patients' therapeutic needs), experienced safety (perceived levels of threat from violence and aggression), and patient cohesion (the perceived presence of peer support between patients). The EssenCes has well established psychometric properties; the factor structure has been replicated in various settings, each factor demonstrates good internal consistency, and construct validity has been demonstrated in tests of convergent and discriminant validity (Howells et al., 2009, Milsom et al., 2014, Tonkin et al., 2012). The 15 scale items are framed by two additional, unscored items (an ice breaker and a concluding item) measured on a unipolar, five-point Likert scale.

The EssenCES was designed to measure the ward climate in forensic mental health settings, but has been used in prisons and learning disability settings (Day et al., 2012, Willets et al., 2014). Violence prevention measures are included in the therapeutic hold and patient cohesion factors. Internal consistency of each factor is good, as measured by Cronbach's alpha (range 0.73 to 0.87; Schalast et al., 2008).

#### *2.4.1.2. Attitudes to Containment Measures Questionnaire (ACMQ; Bowers et al., 2004)*

The ACMQ assesses the views on the acceptability of 11 containment methods for disturbed behaviour (Bowers et al., 2004). The ACMQ lists each of the containment methods with a brief description and a picture, and asks respondents to rate the acceptability of the method using a five-point Likert scale (strongly agree to strongly disagree). Principal components analysis of the scale identified that each method is a separate component (Dack et al., 2012).

## **2.5. Data collection**

For phase 1, two patient focus groups were planned, one with patients from a female ward and one with patients from a male ward. Patient consent was obtained at the start of each group. They were not recorded at the request of patients; one researcher (NH) facilitated the groups and detailed notes were taken by another (GD). Staff interviews were recorded, and conducted and transcribed by a single researcher (NH). Data collection for phases two and three took place from February to October 2014. For the pilot study (phase 2), three wards in one hospital site were purposively chosen to access staff and patients in a variety of settings, a medium secure female ward (the only female ward at this site), and a medium secure and a low secure male ward. For the main study (phase 3) all eligible staff and patients in mental health care pathways were invited to participate. For phases two and three a single researcher (NH) collected all data. Staff were provided with a study questionnaire and a brief verbal explanation of the purpose of the study. Consent was taken to be implied from return of the completed questionnaire. The study was introduced to patients at the weekly ward meetings, and patients identified by the clinical team as having capacity to consent to the study were invited to

participate. A consent form was completed with patient participants, then they were given the option to complete the questionnaire with the researcher, or self-complete it. Most patients (n=77, 81%) chose the former. Participants in the pilot study were also invited to complete the EssenCES and the ACMQ for convergent and discriminant validity testing, and to repeat the scale 7-14 days after the first completion to assess test-retest reliability. With regards to sample size for phase three, 300 participants is sufficient for tool development requiring exploratory factor analysis (Comrey and Lee, 1992); further, Worthington and Whittaker (2006) suggest that a sample size be set prior to data collection. Assuming a 40-item scale, a sample size of 300 falls in the mid-range suggested by Gorsuch (1983), of between 5 to 10 participants per item. Therefore the aim was to recruit a minimum of 300 participants for stage three.

## 2.6. Data analysis

A combination of classical test theory and item response theory (Rasch modelling) was used to assess the psychometric properties of the scale. Statistical analyses were performed using SPSS version 18.0 and Winsteps 3.81.0.

### 2.6.1. Classical test theory

Item test-retest reliability was calculated using weighted kappa with removal of items of less than moderate strength ( $<0.40$ ) based on guidelines provided by Landis & Koch (Landis and Koch, 1977). Descriptive statistics for socio-demographic, employment and clinical characteristics of the participants were produced. Normality of the data were assessed using the Shapiro-Wilk test to ascertain whether to use parametric or non-parametric tests. An exploratory factor analysis (Principal Components Analysis; PCA) was conducted as part of the item reduction process and to evaluate the construct validity of the scale. Internal consistency of the subscales was assessed using Cronbach's alpha coefficients and calculating item-item correlations. Convergent and discriminant validity was measured against the EssenCES and the ACMQ respectively; data were collected from pilot study participants (phase 2) and analysis undertaken once factors of the scale had been identified



(subsequent to phase 3).

### 2.6.2. Item response theory

Unidimensionality of factors identified in the PCA were assessed using principal component analysis of the residuals (PCAr). For comparison, a PCAr was also conducted for the whole scale. The Rasch measurement model is described by the first contrast; an eigenvalue  $\geq 2$  suggests a second dimension (Linacre, 2006,). Further, Rasch assumes that items can be arranged in terms of their 'difficulty' (ratings on 'more difficult' items are lower on the measurement scale while 'easier' items are higher); and individual respondents can be similarly scaled; person-ability, therefore, is the overall level of that tendency across all scale items. 'Variance explained by' refers to the proportion of raw-score variance in the observations explained by item difficulty, person-ability, and rating scale structure (Linacre, 2014) and measures the capacity for the Rasch model to predict the performance of items and persons, with a higher proportion suggesting greater predictability. A 'variance explained by' measure of  $\geq 40\%$  is considered strong,  $\geq 30\%$  moderate, and  $\geq 20\%$  minimal (Conrad et al., 2011). The size of the total unexplained variance is less important than whether the amount of variance is similar to the variance expected if the data fit the Rasch model perfectly, therefore differences between the observed and the expected totals of unexplained variance were also calculated.

Differential item functioning (DIF) can affect fit of the Rasch model to the data (Boone et al., 2014) primarily when the response varies systematically for an individual item between groups despite equal levels of the underlying characteristic (Tennant and Conaghan, 2007). The effect size, DIF contrast (expressed in logits), is the difference between the DIF measures for each group; a 'substantive' DIF is  $\geq 0.64$  logits (Zwick, 1999). Statistical significance was computed using t-tests (Linacre and Wright, 2009). To control for the effects of multiple testing, Bonferroni corrections were made to the t-test  $p$  values (Bland and Altman, 1995). DIF analyses were conducted on all items for gender, for staff/patient group, and between patients completing the scale with and without a researcher. Where DIF is significant, i.e. items with a DIF contrast of  $\geq 0.64$  logits there are various approaches to deal with those items (Linacre, 2015). When deciding how and whether to resolve DIF the trade-off between reliability

and content validity of the measurement scale needs to be considered (Hagquist and Andrich, 2017), therefore decisions were made on an item-by-item basis.

Rasch analysis uses point-measure correlations to investigate whether all scale items measure the same construct. A fundamental concept in Rasch is that higher person measures lead to higher ratings on items and vice versa (Linacre, 2012). The accuracy of this concept is reported by point-measure correlations. All correlations should be noticeably positive ( $>.50$ ).

To measure the item fit to the Rasch model, two statistics are commonly used, the unweighted mean square, and the weighted mean square. An ideal mean square (MNSQ) would be 1.0, where the observed variance is the same as the expected variance and mean square values should be close to 1.0. Values greater than this (underfit) indicate a lack of fit between the items and the model, and values below 1.0 (overfit) indicate that the data predict the model too well, suggesting item redundancy. Validity is more greatly affected by underfit than overfit. The reasonable range of mean square statistic scores for a rating scale is 0.6–1.4 (Wright et al., 1994); items with a MNSQ  $>2.0$  are likely to distort or degrade the scale causing the measurement to be inaccurate, and items with a MNSQ of 1.4 – 2.0 or  $<.5$  are potentially unproductive for the measurement but not degrading (Linacre, 2012).

### **3. RESULTS**

#### **3.1. Scale Development (phase 1)**

Item development was both theory-driven, based on our systematic review of the literature on mental health staff and patients' views about primary and secondary prevention measures (Hallett et al., 2014), and based on local relevant expertise. Two patient focus groups (total  $n=6$ ) and four individual staff interviews were conducted (NH/GD). Patients were current residents on one male ( $n=4$ ) and one female ( $n=2$ ) medium secure ward, while staff were two members of the Prevention and Management of Violence and Aggression (PMVA) training team, the ward manager of a medium secure female ward,

and a qualified, experienced staff nurse working on a medium secure male ward. The data were analysed using thematic analysis, guided by the a priori themes identified in the literature review (Braun and Clarke, 2006). Over 100 initial items were generated; these were reduced to 54 following review by two of the authors. Expert review, by the site research manager and two PhD students at St. Andrew's, and two health-related professors at the University of Northampton,, resulted in removal of 10 items and addition of 10 new items, creating a 54-item scale for pilot testing.

### 3.2. Pilot testing (phase 2)

The pilot scale was completed by 58 staff and 25 patients (overall response rate = 77%). Participants were invited to complete the scale on a second occasion 7-14 days later for test-retest reliability purposes (staff n=36, patients n=5). Application of criteria for item acceptability resulted in retention of 50 items. Linear weighted Kappa was calculated for the 50 remaining statements; 10 statements had fair test-retest reliability (K .21-.40), 37 moderate (K .41-.60) and 3 substantial (K > .61). The 10 statements that showed fair agreement were removed. The final scale with ambiguous, redundant, and non-reliable items removed comprised 40 items. Cronbach's alpha was .87.

### 3.3. Psychometric evaluation (phase 3)

In total, 511 people (352 staff, 144 patients) were invited to participate and 421 completed scales were returned (326 staff, 95 patients), giving an overall response rate of 82% (staff 93%, patients 66%).

Staff and patient characteristics are shown in Table 1.

**Table { SEQ Table \\* ARABIC }. Staff and patient participant characteristics**

Staff	n (%)	Patients	n (%)
<b>GENDER</b>		<b>GENDER</b>	
Male	116 (35.6)	Male	67 (70.5)
Female	170 (52.1)	Female	28 (29.5)
<b>AGE</b>		<b>AGE</b>	
18-24	34 (10.4)	18-24	18 (18.9)
25-34	73 (22.4)	25-34	26 (27.4)

35-44	78 (23.9)	35-44	8 (8.4)
45-54	66 (20.2)	45-54	16 (16.8)
55+	34 (10.4)	55+	15 (15.8)
<b>ROLE</b>		<b>DIAGNOSIS</b>	
Health care assistant	123 (37.7)	Schizophrenia type	29 (30.5)
Qualified nurse	50 (15.3)	Personality disorder	20 (21.1)
Deputy ward manager	11 (3.4)	Personality disorder /	9 (9.5)
Ward manager	4 (1.2)	Schizophrenia type	
Other (occupational therapist, assistant psychologist)	14 (4.3)	Other (inc. developmental, behavioural and mixed)	25 (26.5)
<b>EXPERIENCE</b>		<b>ETHNICITY</b>	
<5 years	100 (30.7)	White	67 (70.5)
5-9 years	32 (9.8)	Black	6 (6.3)
≥10 years	105 (32.2)	Asian	5 (5.3)
		Other	2 (2.1)
<b>TYPE OF WARD</b>		<b>TYPE OF WARD</b>	
Locked	40 (12.3)	Locked	19 (20.0)
Low secure	164 (50.3)	Low secure	38 (40.0)
Medium secure	109 (33.4)	Medium secure	38 (40.0)

### 3.3.1. Principal Components Analysis

The Kaiser-Meyer-Olkin Measure of Sampling Adequacy (.916) suggested that the data were appropriate for PCA. Examination of the scree plot suggested a one-factor solution, whilst findings from the literature review suggested three factors, therefore multiple analyses were run, extracting one, two and three factors. To identify which solution had best fit, two quantitative criteria were used: i) no crossloading items defined as items that loaded .32 or higher on more than one factor, and ii) no factors with fewer than three items (Costello and Osborne, 2005). For the one-factor solution, orthogonal (varimax) rotation was used as the factor correlation was less than .32, Factor correlations for the two- and three-factor solutions exceeded .32, so oblique (direct oblmin) rotation was used in

these cases. For each factor solution, items with a factor loading  $<.50$  were removed before PCA was re-run.

Only the 1- and 2- factor solutions comprised more than three items per factor. The one-factor solution comprised items relating to staff actions and none relating to the actions that patients might take to prevent violence. On inspection, the two-factor solution comprised a staff actions and a patient actions factor. The one-factor solution explained 54.1% of the total variance, slightly higher than the total variance explained by the two-factor solution (53.3%) and so the one-factor solution was marginally statistically stronger than the two-factor solution. However, the loss of all items relating to patient actions in the one-factor solution means that it is conceptually weaker. For this reason, the two-factor scale was chosen for further analysis: staff actions (11 items) and patient actions (5 items). Factor loadings, communalities, and the variance explained by the factors are shown in { [HYPERLINK \l "Ref498085142"](#) }.

**Table 2. Factor loadings with variance**

Item	Loading	Communalities
<b>Factor 1. Staff actions (40.2% variance)</b>		
The staff here are experienced in preventing aggression	.83	.68
Staff here know how to talk to patients	.81	.67
Staff on this ward show the patients respect	.80	.65
Negotiation with aggressive patients is used effectively by staff	.75	.53
The staff have a positive attitude towards patients	.75	.62
Staff know when to intervene when a patient is becoming aggressive	.74	.52
Staff on the ward are good at talking down aggressive patients	.73	.50
Staff are good at listening to patients	.72	.54
Staff here have a good knowledge of the patients	.71	.53
There is usually a member of staff around for patients to talk to	.65	.47
Staff are sometimes rude to patients	.60	.40
<b>Factor 2. Patient actions (13.1% variance)</b>		
Patients bully other patients	.72	.52
Patients are nice to each other	.71	.57
Patients on this ward show the staff respect	.69	.50
Patients on the ward are good at controlling their inner feelings	.66	.43
Patients sometimes annoy other patients on purpose	.66	.42

### 3.3.2. Rasch analysis

PCAr of the 16-item scale resulting from principal components analysis revealed that the eigenvalue of the first contrast for the VPC scale total was 3.6, which verifies scale multidimensionality. The eigenvalue of the first contrast for the staff actions and patient actions factors were below 2.0 (both 1.7), suggesting these factors are unidimensional.

The raw variance explained by measures for the factors was 43.3% (staff actions) and 48.8% (patient actions) and were therefore both considered strong measurement dimensions (Conrad et al., 2011). Comparison of male and female participants revealed no significant item invariance for any items, with all contrasts falling below the 0.64 logit criterion. Comparison of patient and staff participants revealed significant DIF on a single item in the staff actions factor: 'Staff are rude to patients' (DIF contrast .72  $p<.0001$ ). Because this was the only negatively worded staff actions factor item remaining, asking an important question relating to staff behaviour, the first option for dealing with DIF was chosen, to ignore it as inherent in the measurement system and retain the item.

The point-measure correlations for all items were  $\geq .61$ , suggesting high construct validity ({{ HYPERLINK \l "Ref498085041" }}). The infit and outfit MNSQs of all but two items fell within the acceptable range (0.6-1.4) ({{ HYPERLINK \l "Ref4980850411" }}). One item from the patient actions factor had an infit MNSQ of 1.89; given that this fit statistic is unlikely to degrade the scale (i.e., none  $>2.0$ ) it was retained. The staff actions item 'The staff are rude to patients' had an infit MNSQ above 2.0 (2.29); the conceptual importance of the item outweighs the marginal statistical gains made by item removal and so the item was retained.

**Table 3. Item measure and fit for the VPC scale**

Item	Pt Measure	Infit	Infit	Outfit	Outfit Corr. Expected	MNSQ	ZSTD	MNSQ	ZSTD
Staff are good at listening to patients				0.65	0.65	0.98	-0.1	0.99	0.0
There is usually a member of staff around				0.65	0.66	1.0	0.0	1.01	0.2
for patients to talk to									
Staff on the ward are good at talking down				0.60	0.66	1.32	2.3	1.37	2.9
aggressive patients									
The staff have a positive attitude towards				0.65	0.66	0.99	-0.1	0.92	-0.7

patients

**Staff here have a good knowledge of the**

0.69      0.65      0.65      -2.9      0.67      -3.3

**patients**

**Staff on this ward show the patients**

0.65      0.64      0.91      -0.7      0.85      -1.5

**respect**

**Staff know when to intervene when a**

0.66      0.64      0.60      -3.5      0.72      -2.8

**patient is becoming aggressive**

**The staff here are experienced in**

0.70      0.65      .98      -0.1      1.02      0.3

**preventing aggression**

**The staff are rude to patients**

0.63      0.68      2.29      9.9      2.50      9.9

**Negotiation with aggressive patients in**

0.66      0.66      1.26      1.9      1.17      1.5

**used effectively by staff**

**Staff here know how to talk to patients**

0.73      0.65      0.74      -2.1      0.71      -2.9

**Patients are nice to each other**

0.64      0.67      1.38      3.0      2.23      9.9

**Patients bully other patients**

0.68      0.65      1.20      2.0      1.13      1.6

**Patients annoy other patients on purpose**

0.61      0.62      1.89      8.1      1.47      5.8

**Patients on the ward are good at**

0.65      0.67      1.05      0.5      0.94      -0.7

**controlling their inner feelings**

**Patients on this ward show the staff**

0.70      0.67      0.77      -2.2      0.78      -2.4

**respect**

---

Items in bold form the final VPC-14



### 3.3.3. Internal consistency of factors

Cronbach's alpha for the staff actions factor was .92 suggesting redundant items. Item-item correlations were examined and three items were highly correlated ( $>.65$ ) therefore two were removed. The retained item 'Staff on this ward show the patients respect' was retained as it reflects a patient actions factor item and thus retention improved balance. After removal of these items alpha was .89. Alpha for the patient actions factor was .76 suggesting good internal consistency for each factor, with no redundant items.

### 3.3.4. Convergent and discriminant validity

As the data were not normally distributed, the non-parametric Spearman's rank-order test was employed. As predicted, moderate significant correlations were observed between the VPC-14 staff actions factor and the EssenCES therapeutic hold factor,  $r_s(65)=.414$ ,  $p=.001$ , and between the VPC14 patient actions factor and the EssenCES patient cohesion factor,  $r_s(62)=.316$ ,  $p=.012$ . Moderate significant correlations were also observed between the VPC-14 patient actions factor, and the therapeutic hold and experienced safety factors of the EssenCES. Spearman's rank-order correlations between sum factor scores of the VPC scale and the ACMQ ranged from .087 to .223, all weak correlations, and none were significant, thus suggesting good discriminant validity.

## 4. DISCUSSION

The final 14-item version of the scale, the VPC-14, is quick and easy to complete for both staff and patients and was developed using a strong theoretical evidence base. Statistical analysis suggests that it is valid and reliable. All final items were clear and understandable for staff and patients and demonstrated moderate test-retest reliability ( $K .40-.60$ ). PCA indicated a two-factor solution comprising staff actions (F1) and patient actions (F2), with all items having a factor loading  $>.60$  and communalities  $\geq .40$ , and no cross-loading items. Rasch analysis confirmed the unidimensionality of

each factor and suggested good construct validity. Both factors showed good internal consistency, and convergent/discriminant validity was demonstrated.

Despite the initial item pool containing many items relating to organisational and environmental factors, none of these items were retained by the principal components analysis. The antecedents to violence and aggression are more commonly cited as patient or staff factors than organisational or environmental factors (Bowers et al., 2011). It may be that organisational and environmental factors are of greater interest to researchers than participants; these factors frequently appear in quantitative studies of violence prevention, but are rarely found in qualitative studies (Hallett et al., 2014).

#### **4.1. Strengths and limitations**

The strengths of the VPC-14 include its strong theoretical base and developed with expert review and a substantial pilot study with staff and patients. For the psychometric testing, there was a large sample size with excellent response rates. There was a difference in participation rate between staff and patients, 93% and 66% respectively. However this is a particularly high response rate for staff and a good response rate from patients (Aiken et al., 2012) and so the discrepancy is not cause for concern. Item reduction and psychometric testing used a combination of classical test theory and item response theory, thus strengthening statistical confidence in the measure. The use of factor analysis alone can produce varying results between studies, but in combination with Rasch analysis provides greater stability (Kelly et al., 2007). However, test-retest reliability and convergent/discriminant validity were only assessed with a subsample of the participants. Reliability and validity were assessed in accordance with the COSMIN checklist (Mokkink et al., 2010), but responsiveness and interpretability were not. However, responsiveness is about change over time and would require a longitudinal study to assess. Interpretability explores the distribution of scores among groups, and therefore was beyond the scope of this study.

The scale was developed within a single charitable trust, with most participants based in secure settings, therefore it is unclear whether the statistical results would be replicated in other settings.

Including data from a pilot study for further testing with a larger dataset has been used and justified for scale development (Duxbury, 2003). However some researchers argue that responses to items are dependent on their context within a scale (Schwarz, 1999, Stanton et al., 2002).

## **4.2. Implications for practice**

Violence prevention and the reduction of coercive measures (e.g. restraint and seclusion) are receiving much attention currently at organisational and ward level. Wards and trusts in many countries are adopting Safewards, a programme of ten ward-level interventions designed to increase safety, and reduce conflict and containment (Safewards, 2017). Similarly the Six Core Strategies, organisational level interventions to reduce the use of seclusion and restraint, has had widespread adoption in the USA and other countries (Goulet et al., 2017). As the VPC-14 is a measure of staff's and patients' perceptions of the violence prevention climate, it is a useful tool to assess the efficacy of violence or restraint reduction programmes. Even when interventions havenot been introduced, it may be able to provide a meaningful measure of the violence prevention climate over time. It is quick and easy to administer, and can be given to staff and patients, meaning that it can be used to compare the views of these groups.

## **4.3. Further research**

The VPC-14 needs to be evaluated in other mental health settings including general acute to determine the transferability of the psychometric properties. This would also allow confirmatory factor analysis to be undertaken to confirm the factor structure identified by the PCA and Rasch modelling. Finally, this study has not explored differences in the perceptions of the violence prevention climate between and within staff and patient groups.

## 5. CONCLUSION

The VPC-14, based on the knowledge of the authors and considering its psychometric properties, is currently the most robust measure of the inpatient violence prevention climate available. It is quick and easy to administer and can be used with staff and patients, meaning that it has the potential to be introduced as standard practice in a ward setting. It can be used to assess the violence prevention climate over time, as well as to assess the efficacy of new policies and procedures, at least from the perspectives of staff and patients. Using the VPC-14 allows staff and patients to identify elements of the violence prevention climate that they perceive are good or poor, so that the organisation can explore what is working, and also areas that may need improvement. It is hypothesised that improvements to the violence prevention climate would lead to improved staff and patient outcomes, as well as a reduction in conflict and containment events, and so any measure that can help identify where improvements can be made, as well as the efficacy of those improvements once implemented, can only be of benefit to individual staff and patients, and the organisation as a whole.

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